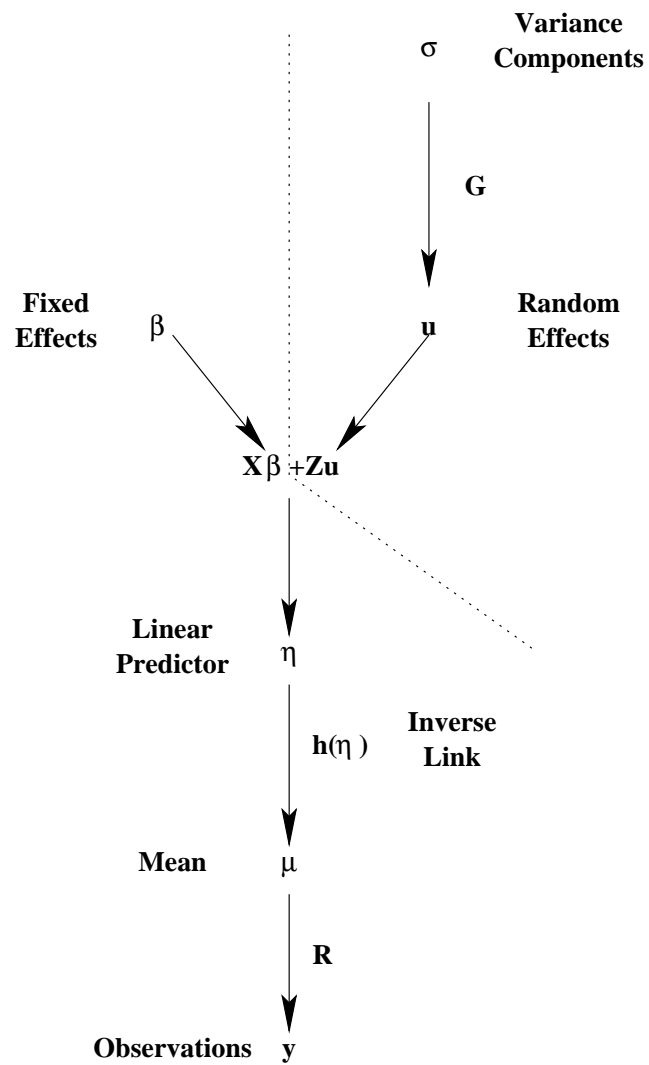


Generalized Linear Mixed Models: Foundations

- Known variance components
- Logistic model
- Marginal likelihood
- Unknown variance components
 - Approximate marginal likelihood
 - MCMC



Density

- Start with $\mathbf{y}|\mathbf{u}$

$$f_Y(\mathbf{y}|\mathbf{u}; \boldsymbol{\beta}, \phi) = \exp \left\{ \sum_{i=1}^N \ell_i \right\}$$

- and the distribution of \mathbf{u}

$$f_U(\mathbf{u}; \boldsymbol{\sigma}) \sim \exp \left\{ \frac{-1}{2} (\ln |D| + \mathbf{u}' D^{-1} \mathbf{u}) \right\}$$

- Get the joint density of \mathbf{y}, \mathbf{u}

$$f_{Y,U}(\mathbf{y}, \mathbf{u}; \boldsymbol{\beta}, \phi, \boldsymbol{\sigma}) = \exp \left\{ \sum_{i=1}^N \ell_i - \frac{1}{2}(\ln |D| + \mathbf{u}' \mathbf{D}^{-1} \mathbf{u}) \right\}$$

- Get the density of \mathbf{y}

$$f_Y(\mathbf{y}; \boldsymbol{\beta}, \phi, \boldsymbol{\sigma}) = \int \dots \int_{\mathbb{R}^q} \exp \left\{ \sum_{i=1}^N \ell_i - \frac{1}{2}(\ln |D| + \mathbf{u}' \mathbf{D}^{-1} \mathbf{u}) \right\} d\mathbf{u}$$

- Difficult to evaluate analytically
 - Numerically
 - Approximately

Penalized Quasi-Likelihood

$$h(\mathbf{u}) = \sum_{i=1}^N \ell_i - \frac{1}{2}(\ln |D| + \mathbf{u}' \mathbf{D}^{-1} \mathbf{u})$$

- Obtaining the the marginal density of \mathbf{y} would be easy if $h(\mathbf{u})$ was a quadratic form.
- Second order Taylor's Series expansion of $h(\mathbf{u})$ about \mathbf{u}_0

$$h(\mathbf{u}) \doteq h(\mathbf{u}_0) + \frac{\partial h}{\partial \mathbf{u}'}(\mathbf{u} - \mathbf{u}_0) + \frac{1}{2}(\mathbf{u} - \mathbf{u}_0)' \frac{\partial^2 h}{\partial \mathbf{u} \partial \mathbf{u}'}(\mathbf{u} - \mathbf{u}_0)$$

with the derivatives evaluated at \mathbf{u}_0 .

- Select \mathbf{u}_0 such that

$$\frac{\partial h}{\partial \mathbf{u}'} \Big|_{\mathbf{u}=\mathbf{u}_0} = 0$$

$$\frac{\partial h}{\partial \mathbf{u}'} = \mathbf{Z}' \mathbf{H}' \mathbf{R}^{-1} (\mathbf{y} - \boldsymbol{\mu}) - \mathbf{D}^{-1} \mathbf{u}$$

- Which yields

$$h(\mathbf{u}) \doteq h(\mathbf{u}_0) + \frac{1}{2} (\mathbf{u} - \mathbf{u}_0)' \frac{\partial^2 h}{\partial \mathbf{u} \partial \mathbf{u}'} (\mathbf{u} - \mathbf{u}_0)$$

- The second partial

$$\frac{\partial^2 h}{\partial \mathbf{u} \partial \mathbf{u}'} = -\mathbf{Z}' \mathbf{H}' \mathbf{R}^{-1} \mathbf{H} \mathbf{Z} - \mathbf{D}^{-1}$$

- Which yields

$$h(\mathbf{u}) \doteq \sum_{i=1}^N \ell_i(\mathbf{u}_0) - \frac{1}{2}(\ln |D| + \mathbf{u}_0' \mathbf{D}^{-1} \mathbf{u}_0) - \frac{1}{2}(\mathbf{u} - \mathbf{u}_0)' [\mathbf{Z}' \mathbf{H}' \mathbf{R}^{-1} \mathbf{H} \mathbf{Z} + \mathbf{D}^{-1}] (\mathbf{u} - \mathbf{u}_0)$$

- Integrating yields

$$\ell \doteq \sum_{i=1}^N \ell_i(\mathbf{u}_0) - \frac{1}{2}(\ln |D| + \mathbf{u}_0' \mathbf{D}^{-1} \mathbf{u}_0) - \frac{1}{2} \ln |\mathbf{Z}' \mathbf{H}' \mathbf{R}^{-1} \mathbf{H} \mathbf{Z} + \mathbf{D}^{-1}|$$

- Taking the partial with respect β

$$\frac{\partial \ell}{\partial \beta} \doteq \mathbf{X}' \mathbf{H}' \mathbf{R}^{-1} (\mathbf{y} - \boldsymbol{\mu})$$

- Approximate MLE by solving

$$\begin{aligned} \mathbf{X}' \mathbf{H}' \mathbf{R}^{-1} (\mathbf{y} - \boldsymbol{\mu}) &= \mathbf{0} \\ \mathbf{Z}' \mathbf{H}' \mathbf{R}^{-1} (\mathbf{y} - \boldsymbol{\mu}) - \mathbf{D}^{-1} \mathbf{u} &= \mathbf{0} \end{aligned}$$

- Using Fisher scoring

$$\begin{pmatrix} X' H' R^{-1} H X & X' H' R^{-1} H Z \\ Z' H' R^{-1} H X & Z' H' R^{-1} H Z + D^{-1} \end{pmatrix} \begin{pmatrix} \hat{\beta} \\ \hat{u} \end{pmatrix} = \begin{pmatrix} X' H' R^{-1} (\mathbf{y} - \mu + H\eta) \\ Z' H' R^{-1} (\mathbf{y} - \mu + H\eta) \end{pmatrix}$$

- Has the advantage of being simple to implement.
- Depends on how closely the conditional distribution of \mathbf{y} given \mathbf{u} approximates a normal distribution.
 - Work well **when** the central limit theorem applies

Estimation of variance components

- Recall that REML in the linear mixed model case involved setting a set of quadratic forms equal to their expectation and solving.

$$\hat{\mathbf{u}}_i' \hat{\mathbf{u}}_i$$

- Using the same set of quadratic forms and their approximate expectation:

$$\begin{pmatrix} \hat{\beta} \\ \hat{u} \end{pmatrix} = C \begin{pmatrix} X' H' R^{-1} \\ Z' H' R^{-1} \end{pmatrix} (y - \hat{\mu} + H\hat{\eta})$$

$$(y - \mu + H\eta) \sim (HX\beta, R + HZDZ'H')$$

$$E \begin{pmatrix} \hat{\beta} \\ \hat{u} \end{pmatrix} \doteq C \begin{pmatrix} X' H' R^{-1} HX\beta \\ Z' H' R^{-1} HX\beta \end{pmatrix}$$

$$E(\hat{\beta}) \doteq [C^{XX} X' H' R^{-1} HX + C^{XZ} Z' H' R^{-1} HX] \beta$$

$$E(\hat{u}) \doteq 0$$

$$\begin{aligned}
\text{var} \begin{pmatrix} \hat{\beta} \\ \hat{u} \end{pmatrix} &= C \begin{pmatrix} X'H'R^{-1}HX & X'H'R^{-1}HZ \\ Z'H'R^{-1}HX & Z'H'R^{-1}HZ \end{pmatrix} C \\
&\quad + C \begin{pmatrix} X'H'R^{-1}HZ \\ Z'H'R^{-1}HZ \end{pmatrix} D \begin{pmatrix} Z'H'R^{-1}HX \\ Z'H'R^{-1}HZ \end{pmatrix} C \\
&= [C - CD^{-1}C] + [D - C - C + CD^{-1}C] \\
\hat{u} &\sim (0, D - C^{ZZ})
\end{aligned}$$

Which is identical to the linear mixed model case.

- The estimating equations become

$$.5 \left[\text{Diag} \left(\frac{1}{\sigma_i^4} \{q_i - 2 \text{tr}(C^{ii}) / \sigma_i^2\} \right) + \left\{ \frac{1}{\sigma_i^4 \sigma_j^4} \text{tr}(C^{ij} C^{ji}) \right\} \right] \hat{\sigma} = .5 \left\{ \frac{\hat{u}'_i \hat{u}_i}{\sigma_i^4} \right\}$$

Germination Rate

- Data⁵ from an experiment examining the germination rates
- Two types of seeds (*O. aegyptiaco* 75 and 73)
- Two types of root extracts (bean and cucumber).
- Dependent variable: Number of seeds that germinated r_{ijk} out of the the total number present in a dish n_{ijk} . table.

⁵Source: Crowder, M. (1978) *Beta-binomial Anova for proportions*. Appl. Statist. **27**, 34–37 (Table 3)

<i>O. aegyptiaco</i> 75				<i>O. aegyptiaco</i> 73			
Bean		Cucumber		Bean		Cucumber	
r_{11k}	n_{11k}	r_{12k}	n_{12k}	r_{21k}	n_{21k}	r_{22k}	n_{22k}
10	39	5	6	8	16	3	12
23	62	53	74	10	30	22	41
23	81	55	72	8	28	15	30
17	39	46	79	0	4	3	7
		10	13				

$$y_{ijk} = r_{ijk}/n_{ijk} | \mathbf{u} \sim \frac{1}{n_{ijk}} \text{Bin}(n_{ijk}, \mu_{ijk})$$

$$\mathbb{E}(y_{ijk} | \mathbf{u}) = \mu_{ijk}$$

$$\text{var}(y_{ijk} | \mathbf{u}) = \frac{\mu_{ijk}(1 - \mu_{ijk})}{n_{ijk}}$$

$$\ell(\boldsymbol{\mu}; \mathbf{y} | \mathbf{u}) = c + \sum_{i=1}^2 \sum_{j=1}^7 n_{ijk} [y_{ijk} \ln(\mu_{ijk}) + (1 - y_{ijk}) \ln(1 - \mu_{ijk})]$$

The linear predictor is

$$\eta_{ij} = \mu + T_i + E_j + (TE)_{ij} + D_{ijk}$$

- μ is the intercept
- T_i is the main effect of seed type i
- E_j is the main effect of root extract j ,
- D_{ijk} : is a random dish effect iid $N(0, \sigma_D^2)$.

ASREML:

Germination Example

```
type !A
extract !A
r
n
y
plate !I
germ.dat !MAXIT=100
r !BINOMIAL !TOTAL n ~ type*extract !r plate
0 0 1
plate 1
0 0 I 1
predict type extract !TDIFF
predict type !TDIFF
```

Results:

1	LogL=-6.22355	S2= 1.0000	13 df	1.000	1.000
2	LogL=-5.92972	S2= 1.0000	13 df	1.000	0.9524
3	LogL=-3.76174	S2= 1.0000	13 df	:	1 components constrained
4	LogL=-.878734	S2= 1.0000	13 df	1.000	0.3252E-01
5	LogL=-.811204	S2= 1.0000	13 df	1.000	0.4485E-01
6	LogL=-.787669	S2= 1.0000	13 df	1.000	0.5803E-01
7	LogL=-.787653	S2= 1.0000	13 df	1.000	0.5904E-01
8	LogL=-.787651	S2= 1.0000	13 df	1.000	0.5901E-01
Final parameter values				1.0000	0.59009E-01
Deviance from GLM fit		13	10.55		
Variance heterogeneity factor [Deviance/DF]			0.81		

Degrees of Freedom and Stratum Variances

1.35 0.590093E-01 1.0

Source	Model terms	Gamma	Component	Comp/SE	% C
plate	identity	17	0.590093E-01	0.590093E-01	0.82 0 U

Analysis of Variance	NumDF	DenDF	F_inc	Prob
1 type	2	10.8	2.35	0.141
2 extract	1	9.7	30.96	<.001
7 type.extract	1	13.0	4.54	0.053

	Estimate	Standard Error	T-value	T-prev
7 type.extract				
4	-0.978362	0.460229	-2.13	
2 extract				
Cucumber	1.54346	0.267891	5.76	
1 type				
0_aeg_75	-0.695085	0.189894	-3.66	
0_aeg_73	-0.696893	0.276081	-2.52	-0.01
6 plate				
		17 effects fitted		